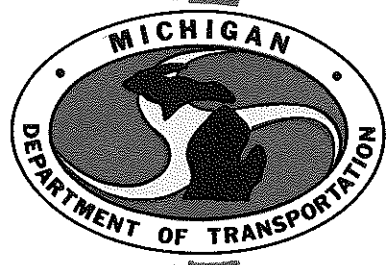


PRECISIONS OF THE AGGREGATE SAMPLE
SPLITTER AND TESTING METHOD



**TESTING AND RESEARCH DIVISION
RESEARCH LABORATORY SECTION**

PRECISIONS OF THE AGGREGATE SAMPLE
SPLITTER AND TESTING METHOD

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Michigan Transportation Commission
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SUMMARY

Aggregate samples obtained from sources for inspection purposes are usually greater than the capacity of the testing equipment; thus, samples are generally reduced before testing. The usual practice is to divide the sample into subsamples with a Gilson splitter. However, only one of the two subsamples is tested for inspection purposes. Of concern is whether the splitter is capable of consistently splitting a sample into two subsamples of equal gradation.

Also, the Department is currently implementing an in-place aggregate inspection procedure. The aggregate samples are to be tested using a new mechanical testing method. It is of interest to know whether or not the 'hand' and mechanical testing methods produce the same gradation results.

The following conclusions are drawn from three experiments described in this report:

- 1) The sample splitter is able to split a sample of aggregate into two subsamples such that their weight ratio is very nearly constant. However, this constant may not be the ratio of the number of openings on the two sides of the splitter.

- 2) The aggregate gradations of the two subsamples are slightly different. The differences, however, are negligible relative to in-place aggregate uniformity. Hence, either subsample can be used to represent the original sample for in-place aggregate inspection purposes.

- 3) The testing precision of the hand and mechanical methods are very high relative to in-place aggregate uniformity.

- 4) The aggregate gradation differences measured by the two testing methods are negligible based on laboratory samples, relative to the in-place aggregate uniformity. The differences, however, are significant based on field samples. After reviewing laboratory and field sample preparation procedures, it was concluded that the mechanical testing method degrades aggregate more than the hand testing method. Intuitively, the degree of degradation depends on the testing time. This suggests studying the degradation time profile of the mechanical testing method for the purpose of properly adjusting aggregate specifications.

INTRODUCTION

Aggregate samples obtained from construction sites for inspection purposes are usually greater than the capacity of the testing equipment; thus, samples have to be reduced in size before testing. The current practice is to use a Gilson sample splitter to divide a sample into two subsamples. Only one of the two subsamples is tested using the hand or mechanical testing method. The test results are then used to accept or reject the in-place aggregate according to an acceptance sampling plan. This practice would not be meaningful unless the following questions can be answered in the affirmative.

- 1) Is the sample splitter capable of consistently splitting a sample into two subsamples of equal gradation?
- 2) Do the hand and mechanical testing methods produce the same gradation results?
- 3) Are splitting and testing precisions high, relative to in-place aggregate uniformity?

The major purpose of this research study is to answer the above questions in statistical terms. For this purpose, we conducted three experiments taking samples of 22A aggregate from construction projects currently inspected by a newly developed in-place aggregate sampling plan.¹ These experiments were designed so that normal inspection practice would not be interrupted. Before describing the experiments, we shall define several key terms used throughout the text.

'Testing error' is defined as the difference between the true value and the test results. 'Testing precision' is, by definition, the variance of the testing error. Similarly, 'splitting error' (on the subsample) is defined as the difference of the true values of the sample and the subsample. 'Splitting precision' is, therefore, the variance of the splitting error. By like reasoning, 'in-place aggregate uniformity' can be defined as the variance of the in-place aggregate gradation. Notice that high testing and splitting precisions as well as high in-place aggregate uniformity imply low testing, splitting, and aggregate variance.

¹ Kuo, Wen-Hou, "Statistical Analysis of Aggregate Base Course Inspection Using an End Result Aggregate Specification," Michigan Department of Transportation, Research Report No. R-1024, February 1977.

TABLE 1
AGGREGATE GRAIN SIZE SPECIFICATIONS

Pile No.	Aggregate Grain Size
1	Passes 1-in. sieve, and is retained on 3/4-in. sieve
2	Passes 3/4-in. sieve, and is retained on 3/8-in. sieve
3	Passes 3/8-in. sieve, and is retained on No. 4 sieve
4	Passes No. 4 sieve, and is retained on No. 8 sieve
5	Passes No. 8 sieve

TABLE 2
TARGET AGGREGATE GRADATIONS
FOR MAKING UP GROUP SAMPLES

Group No.	Percent Passing Sieves			
	3/4-in.	3/8-in.	No. 4	No. 8
1	98	67	57	32
2	98	69	59	34
3	98	71	61	36
4	98	73	63	38
5	98	75	65	40
6	98	77	67	42
7	98	79	69	44
8	98	81	71	46
9	98	83	73	48
10	98	85	75	50

EXPERIMENTAL PROCEDURES

Experiment A

This experiment was designed to investigate testing precision using laboratory samples. The data preparation is described below.

Aggregates obtained from Lot No. 2 of Construction Project FR 23092-10729A² were separated into five piles of different aggregate sizes specified in Table 1. Ten sets of four samples were then made up from these five piles according to the predetermined target gradations specified in Table 2. Each set of four consisted of two 2,000-gm and two 4,000-gm samples. One of the two samples in each weight group was tested using the hand testing method and the other was tested using the mechanical testing method, the same operator performing all tests. ASTM Procedure C136 governed the manual shaking. For the mechanical shaking, the 4,000-gm and 2,000-gm samples were shaken approximately 11 and 8 minutes, respectively. Loss-by-hand washing was determined in accordance with ASTM C117, and by the mechanical washing apparatus for the mechanical washing. The end points of the wash were determined as specified in the method. The test results of 40 samples for this experiment are presented in Table 3.

Experiment B

This experiment was designed, using the field samples, to investigate the splitting precision of the sample splitter. Thirty-six samples obtained from Lots 4 through 6 of Construction Project M 36021-10139A³ were processed as follows: a Gilson sample splitter was used to reduce samples to smaller size. This splitter has four openings on one side and five openings on the other side. The sides with four openings and five openings were designated Side 1 and Side 2, respectively. Each sample of 40 to 50 lb was put through the splitter. The material from Side 2 was discarded. The weight of the material from Side 1 averaged 9,870 gm. This portion of the material was put through the splitter again to produce the two final samples. Material from Side 1 (Subsample 1) was used for the inspection purpose. Material from Side 2 (Subsample 2) was used, together with Subsample 1, to study the splitting precision of the sample splitter. Both subsamples were tested using a mechanical shaker and washing machine (mechanical

² This is one of the four construction projects inspected by the newly developed in-place aggregate acceptance sampling plan.

³ Ibid.

TABLE 3
TEST RESULTS OF SAMPLES PREPARED
FOR EXPERIMENT A

Group No.	Sample Weight, gm	Testing Method	Percent Passing Sieves				Percent Loss-by-Washing
			3/4-in.	3/8-in.	No. 4	No. 8	
1	2000	H*	98.00	67.45	56.95	33.55	4.85
	2000	M	98.00	67.50	57.70	33.25	6.00
	4000	H	97.68	67.93	56.95	33.38	5.48
	4000	M	98.03	67.83	57.93	33.70	6.13
2	2000	H	97.65	69.42	59.07	35.63	5.60
	2000	M	97.95	69.45	59.75	35.30	6.25
	4000	H	97.65	70.10	59.23	35.43	5.53
	4000	M	98.25	69.63	59.84	35.52	5.92
3	2000	H	98.00	71.80	61.00	37.35	5.55
	2000	M	98.10	71.17	61.76	37.14	6.56
	4000	H	97.85	71.99	60.94	37.36	5.68
	4000	M	98.00	71.51	61.61	37.24	6.45
4	2000	H	97.95	74.06	62.82	39.18	6.10
	2000	M	96.25	73.51	62.97	39.53	6.60
	4000	H	98.20	73.66	62.86	39.19	6.20
	4000	M	97.72	73.40	63.48	39.40	6.43
5	2000	H	97.60	75.38	65.01	41.14	6.46
	2000	M	98.50	75.35	65.45	41.10	7.00
	4000	H	97.55	75.34	64.91	41.27	6.20
	4000	M	98.00	75.53	65.53	41.20	6.98
6	2000	H	98.00	77.50	66.90	43.15	6.85
	2000	M	98.55	77.20	67.65	43.15	7.25
	4000	H	97.32	77.79	66.84	43.04	6.75
	4000	M	98.05	77.31	67.61	43.30	7.15
7	2000	H	98.05	80.09	68.88	45.17	6.70
	2000	M	98.00	79.44	69.24	45.17	7.20
	4000	H	97.77	79.62	68.71	45.25	7.13
	4000	M	97.77	79.37	69.29	45.06	7.25
8	2000	H	97.00	81.43	70.87	47.75	7.61
	2000	M	98.00	81.10	71.45	46.95	7.20
	4000	H	97.83	81.35	70.53	47.10	7.18
	4000	M	97.70	81.31	71.36	47.25	7.45
9	2000	H	97.55	83.94	72.99	49.63	7.50
	2000	M	98.10	83.19	73.44	48.92	7.50
	4000	H	97.75	83.45	72.54	49.11	6.98
	4000	M	98.03	83.42	72.87	48.99	7.60
10	2000	H	97.40	85.22	74.68	51.40	7.69
	2000	M	97.30	85.18	75.38	50.75	7.61
	4000	H	97.20	85.25	74.55	51.18	7.60
	4000	M	98.03	85.25	75.29	51.19	7.65

* H and M stand for the hand and mechanical testing methods.

testing method). This testing method requires about 20 minutes to wash and 15 minutes to shake. The above procedures are outlined in Figure 1, and the test results are given in Table 4. The percent passing the 1-in. sieve is 100 for every subsample and, therefore, is not included in Table 4.

Experiment C

This experiment was designed to simultaneously study, using the field samples, the splitting and testing precisions. The sample preparation is described below. Eighty-four samples taken from Lots 4 to 10 of Construction Project FR 64015-11535A⁴ were processed as follows: the average weight of these 84 samples is 6,449 gm. Each sample was split into two portions by a sample splitter. This splitter has an equal number of openings on both sides. One side of the splitter was designated as Side 1 and the other side was designated as Side 2. The material from Side 1 and Side 2 was called Subsample I and II, respectively. This is referred to as the first stage splitting. Subsample II was mechanically tested and was used for inspection purposes. Subsample I was further split into two subsamples, Subsample I-1 and Subsample I-2. This is referred to as the second stage splitting. These two subsamples were tested using the hand testing method. The above procedures are outlined in Figure 2, and test results are presented in Table 5. Again, the percent passing the 1-in. sieve is 100 for every subsample and, therefore, not included in the table.

Remarks

We see from the above experiments that testing precisions were investigated using both laboratory and field samples, and that splitting precisions were investigated using two slightly different sample splitters. For discussion purposes, we shall identify the sample splitter used in Experiments B and C as Splitter B and Splitter C, respectively, throughout the text.

The detailed statistical techniques used to analyze the above experiments are presented in Appendices A, B, and C for readers who are interested. The findings of these experiments are presented in the next two sections. To facilitate the discussion of results, we present Table 6 which gives the in-place aggregate uniformities of construction projects known to the author.

⁴Ibid.

TABLE 4
TEST RESULTS OF SAMPLES PREPARED
FOR EXPERIMENT B

Sample No.	Subsample No.	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
			3/4-in.	3/8-in.	No. 8	
1	1	3215	95.89	70.82	49.27	5.69
	2	4878	92.62	63.57	41.78	6.15
2	1	3585	95.20	69.51	48.31	6.42
	2	5160	95.41	67.17	43.99	6.88
3	1	3090	91.49	66.34	46.21	6.54
	2	4563	92.99	65.64	42.67	7.12
4	1	3380	93.28	65.53	45.41	6.07
	2	5483	94.86	65.75	42.49	6.35
5	1	3507	93.36	68.26	48.73	6.19
	2	5310	92.98	65.57	44.26	6.74
6	1	3718	97.31	71.03	49.52	7.15
	2	5090	92.69	66.62	44.81	6.58
7	1	3873	94.14	69.20	48.54	7.64
	2	5398	94.16	65.73	44.74	7.87
8	1	3748	93.76	77.13	47.89	7.10
	2	5370	95.01	65.59	43.99	7.04
9	1	3584	91.63	64.23	43.81	7.17
	2	4888	89.81	62.46	42.04	7.02
10	1	3665	90.67	63.66	43.47	6.82
	2	4698	93.87	64.24	43.27	6.07
11	1	4234	94.80	63.16	40.08	6.05
	2	4728	89.91	63.68	40.84	6.20
12	1	4473	94.86	64.45	43.89	5.72
	2	5515	92.33	65.98	44.13	5.89
13	1	5090	92.99	66.88	46.74	6.35
	2	6438	92.79	65.39	44.80	6.20
14	1	4777	93.34	70.86	51.22	5.80
	2	6623	93.10	75.80	57.69	5.72
15	1	3952	93.52	66.24	46.26	6.00
	2	4835	93.13	63.91	43.12	5.73
16	1	4036	94.33	70.99	48.49	5.77
	2	5515	95.09	71.22	50.08	5.53
17	1	4317	95.39	70.93	50.06	5.37
	2	6350	94.69	69.48	50.39	5.34
18	1	3872	93.29	70.27	49.35	6.40
	2	6512	93.80	69.70	49.49	6.59

TABLE 4 (Cont.)
TEST RESULTS OF SAMPLES PREPARED
FOR EXPERIMENT B

Sample No.	Subsample No.	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
			3/4-in.	3/8-in.	No. 8	
19	1	4691	95.89	71.63	49.31	6.76
	2	6888	93.92	72.71	49.40	7.20
20	1	4709	93.31	68.46	48.23	5.86
	2	7090	93.81	68.41	47.80	5.44
21	1	5268	93.39	72.49	52.60	5.90
	2	6570	95.30	72.50	49.19	6.10
22	1	5281	93.90	71.01	47.93	6.29
	2	6602	92.90	67.49	47.59	5.91
23	1	5004	93.39	70.68	50.48	7.89
	2	6253	91.91	68.11	47.42	7.71
24	1	4613	92.89	70.21	44.81	6.81
	2	5811	93.91	73.02	49.22	6.69
25	1	3947	93.72	72.18	50.39	6.06
	2	5065	95.60	73.39	49.63	6.08
26	1	4154	94.03	71.40	51.54	6.48
	2	5287	95.71	69.62	48.61	5.83
27	1	5237	93.45	69.79	49.74	6.68
	2	6393	94.40	68.48	47.43	6.77
28	1	4587	96.60	71.81	50.77	6.65
	2	6320	94.59	71.14	47.34	5.32
29	1	4378	96.39	72.45	50.59	6.99
	2	5818	94.91	71.52	47.61	6.77
30	1	3808	94.80	67.44	48.08	6.57
	2	5272	93.00	63.66	44.52	5.96
31	1	4953	93.30	68.67	48.68	6.32
	2	6110	92.75	66.68	45.43	5.53
32	1	3803	91.06	64.76	46.52	5.21
	2	5344	93.32	64.84	44.69	5.78
33	1	3907	92.24	70.62	50.96	6.53
	2	6000	96.50	66.37	45.92	6.67
34	1	4248	96.23	70.39	49.91	6.90
	2	6122	93.83	66.11	44.95	7.14
35	1	4245	93.83	70.84	49.33	6.36
	2	5428	93.90	69.10	50.99	6.60
36	1	3565	94.31	70.69	48.81	6.03
	2	5093	95.17	68.98	51.78	6.30

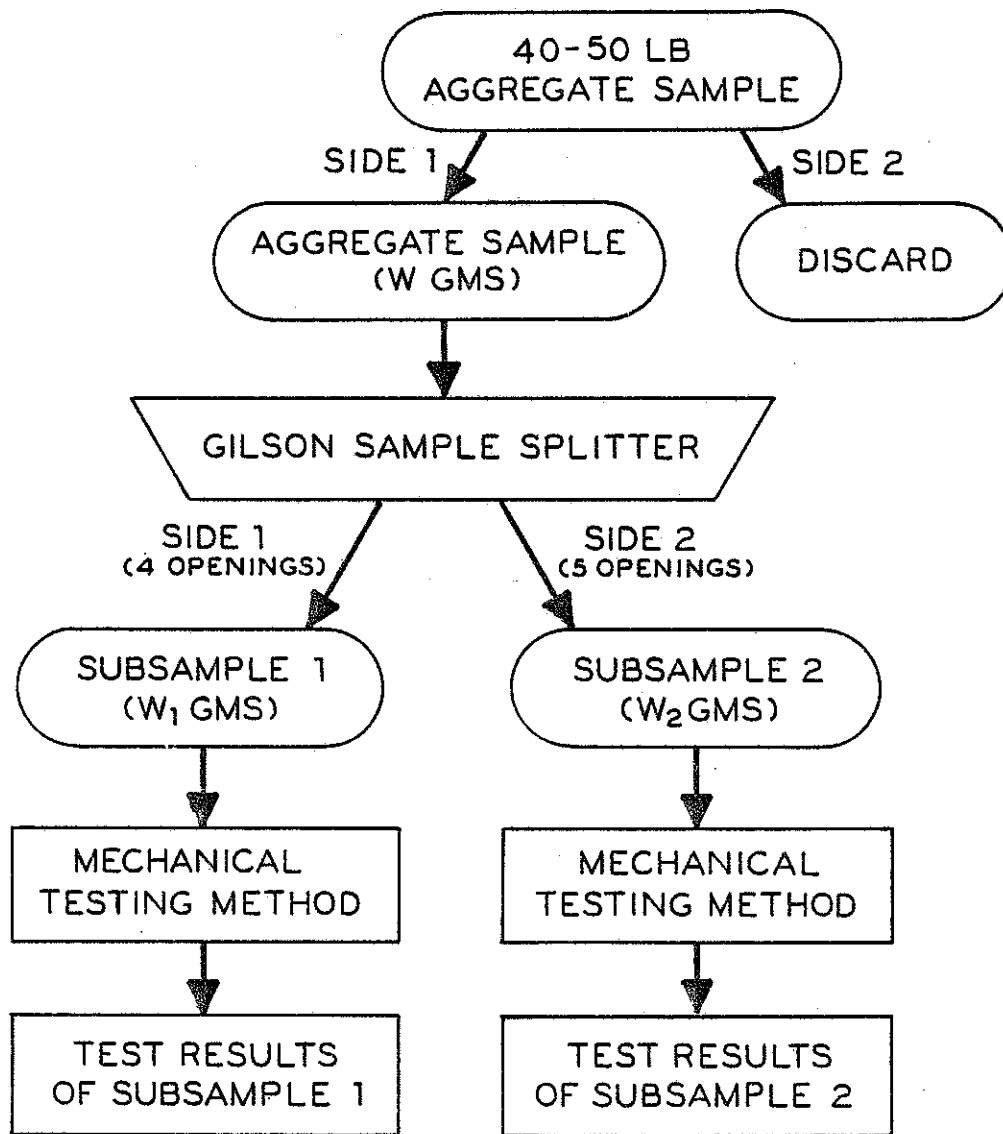


Figure 1. Sample preparation procedures of Experiment B.

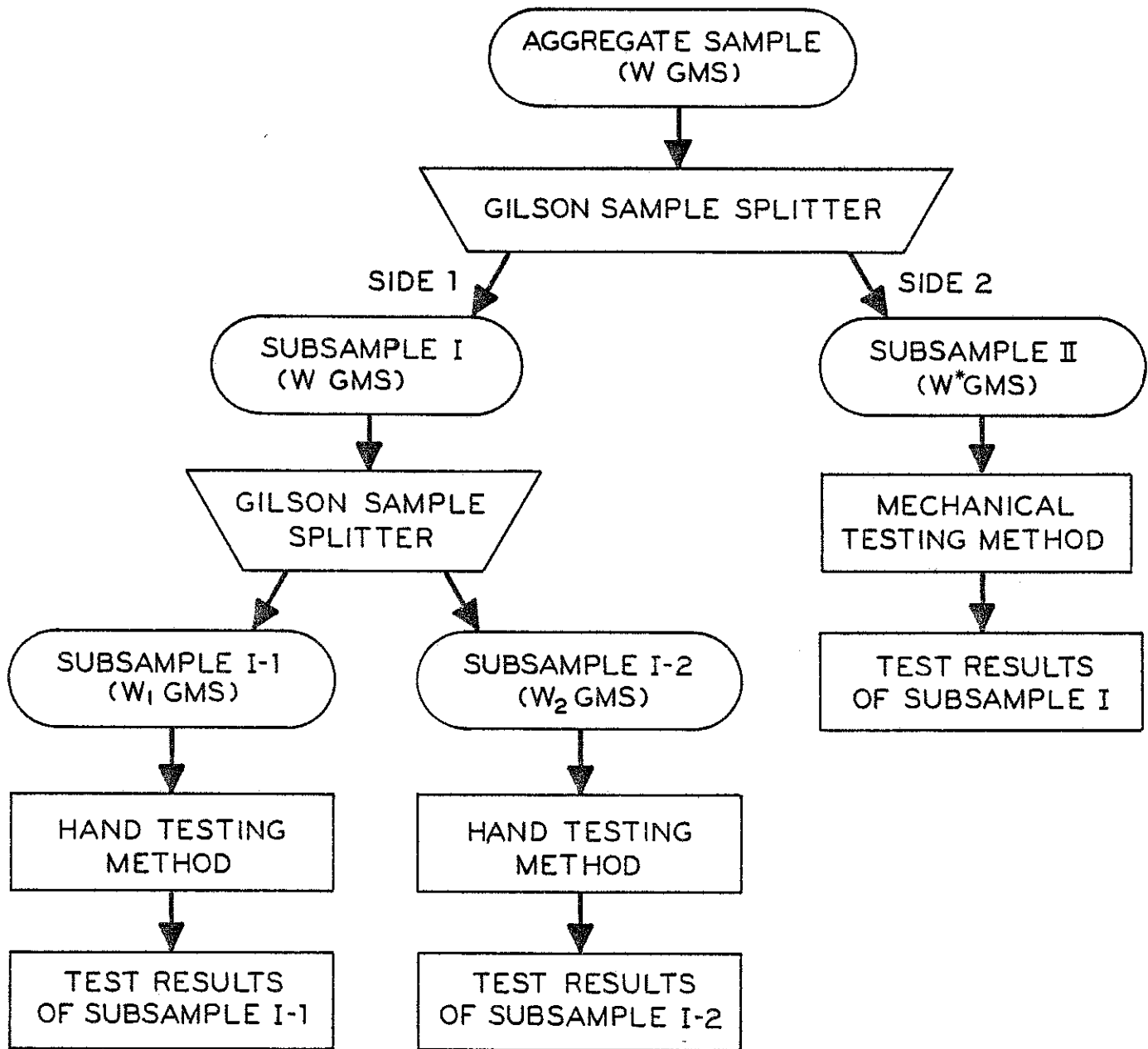


Figure 2. Sample preparation procedures of Experiment C.

TABLE 5
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
1	II	M	3628	95.98	74.31	46.91	8.27
	I-1	H	1481	92.57	69.21	44.63	7.43
	I-2	H	1559	94.10	71.46	43.94	7.63
2	II	M	2762	95.94	73.86	47.25	8.51
	I-1	H	1265	97.23	75.65	45.85	7.51
	I-2	H	1457	97.25	76.39	45.64	7.82
3	II	M	3030	95.58	72.87	46.34	8.09
	I-1	H	1496	91.98	70.92	42.25	7.42
	I-2	H	1274	93.01	71.59	44.58	7.46
4	II	M	3076	96.29	72.11	45.94	8.65
	I-1	H	1416	93.64	66.45	40.25	8.12
	I-2	H	1435	90.31	66.69	40.70	7.39
5	II	M	3157	95.19	75.55	47.51	8.46
	I-1	H	1479	94.25	75.05	45.98	8.25
	I-2	H	1503	96.81	72.32	43.71	8.05
6	II	M	2610	96.55	70.61	43.22	7.70
	I-1	H	1557	92.04	71.10	43.16	7.90
	I-2	H	1337	89.08	67.02	40.61	7.18
7	II	M	2783	95.33	72.58	44.16	7.83
	I-1	H	1570	93.69	67.64	41.40	7.01
	I-1	H	1445	97.16	70.80	41.66	6.71
8	II	M	3018	94.20	72.00	44.76	7.69
	I-1	H	1327	94.80	68.73	42.43	7.01
	I-2	H	1777	92.85	68.37	42.04	7.03
9	II	M	3461	94.19	73.94	47.10	7.37
	I-1	H	1740	94.43	72.87	43.74	6.55
	I-2	H	1584	95.08	72.41	43.43	6.76
10	II	M	3436	97.64	76.89	49.85	6.87
	I-1	H	1602	94.88	75.84	45.51	6.18
	I-2	H	1277	96.01	77.29	47.77	6.11
11	II	M	3736	96.33	73.45	47.78	7.60
	I-1	H	1592	89.82	66.08	39.82	6.03
	I-2	H	1396	94.77	71.92	45.42	6.73
12	II	M	3559	95.62	74.29	46.87	8.12
	I-1	H	1586	94.45	71.37	43.44	7.63
	I-2	H	1658	93.49	71.59	43.85	7.54

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
13	II	M	3525	94.64	72.09	46.67	7.38
	I-1	H	1661	95.06	69.54	42.99	7.10
	I-2	H	1382	95.88	74.02	46.53	7.31
14	II	M	4052	97.21	77.89	50.22	8.02
	I-1	H	1668	96.76	72.78	43.35	7.43
	I-2	H	1465	95.90	78.36	51.54	7.44
15	II	M	2949	96.13	77.08	48.42	7.19
	I-1	H	1385	96.25	71.70	42.45	6.50
	I-2	H	1595	93.86	74.86	44.20	6.83
16	II	M	3433	95.34	75.21	48.06	7.72
	I-1	H	1429	89.78	63.54	37.51	6.86
	I-2	H	1747	90.44	67.60	40.53	7.04
17	II	M	3088	94.30	73.09	47.02	7.58
	I-1	H	1442	91.82	69.83	41.47	6.73
	I-2	H	1655	93.60	69.55	43.08	7.31
18	II	M	3345	98.03	76.47	50.55	7.44
	I-1	H	1237	90.95	72.92	45.43	6.79
	I-2	H	1272	93.79	72.88	47.80	7.31
19	II	M	3596	95.75	76.08	48.47	6.98
	I-1	H	1858	96.77	77.18	47.20	7.16
	I-2	H	2060	94.22	75.00	45.24	7.23
20	I	M	3591	94.96	74.69	47.79	6.99
	I-1	H	1113	91.64	69.09	42.14	6.20
	I-2	H	1414	90.38	70.79	43.71	6.86
21	II	M	3289	96.20	74.10	47.86	7.02
	I-1	H	1359	95.95	74.76	46.21	7.14
	I-2	H	1414	96.96	70.23	43.49	6.58
22	II	M	3245	95.32	72.11	44.84	7.86
	I-1	H	1339	90.74	64.45	39.81	7.02
	I-2	H	1646	98.78	72.72	43.74	7.72
23	II	M	3745	95.41	75.57	47.93	7.08
	I-1	H	1447	90.26	69.25	44.57	6.91
	I-2	H	1905	92.76	70.92	42.73	6.93
24	II	M	3195	94.87	74.74	46.73	7.07
	I-1	H	1343	96.28	73.72	42.81	7.07
	I-2	H	1525	90.36	64.85	38.30	6.49

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
25	II	M	3220	98.88	76.43	48.54	7.92
	I-1	H	1366	92.31	70.06	42.09	7.69
	I-2	H	1479	95.74	74.04	47.19	8.11
26	II	M	3422	95.50	74.69	47.81	7.57
	I-1	H	1562	94.94	71.06	42.45	7.17
	I-2	H	1463	96.31	71.09	45.86	7.38
27	II	M	3277	95.33	74.18	45.50	7.78
	I-1	H	1438	91.52	67.80	41.17	8.83
	I-2	H	1653	90.80	69.21	43.32	7.44
28	II	M	3263	94.09	74.93	48.05	7.85
	I-1	H	1634	94.74	73.19	45.23	7.53
	I-2	H	1568	93.88	72.13	45.34	7.72
29	II	M	3063	98.92	81.42	51.98	7.61
	I-1	H	1874	97.60	77.21	48.40	7.20
	I-2	H	1916	96.71	76.72	47.70	7.05
30	II	M	3192	96.74	75.69	47.34	8.15
	I-1	H	1604	92.46	66.58	42.08	7.36
	I-2	H	1584	90.40	65.66	40.47	7.01
31	II	M	3576	94.71	72.01	45.69	6.77
	I-1	H	1533	96.87	71.43	43.12	6.91
	I-2	H	1671	91.02	69.24	44.46	6.40
32	II	M	3421	96.32	75.56	47.85	7.31
	I-1	H	1543	94.49	68.96	41.87	7.19
	I-2	H	1688	94.19	73.16	46.98	6.99
33	II	M	3805	95.56	73.40	46.65	7.39
	I-1	H	1873	91.35	71.01	43.94	7.21
	I-2	H	1833	92.47	70.32	45.61	6.93
34	II	M	3457	94.82	71.22	45.88	7.06
	I-1	H	1634	94.25	71.42	45.47	6.73
	I-2	H	1662	90.79	69.55	43.50	7.04
35	II	M	3428	94.19	72.37	46.12	7.26
	I-1	H	1964	95.47	70.21	42.11	6.57
	I-2	H	1958	91.52	69.71	43.31	6.69
36	II	M	3342	95.99	74.75	47.85	7.24
	I-1	H	1503	93.68	73.65	45.11	7.45
	I-2	H	1665	97.96	79.16	48.23	6.97

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
37	II	M	3875	96.98	74.81	48.46	7.33
	I-1	H	1448	90.61	71.62	44.34	7.39
	I-2	H	1932	94.62	75.21	46.58	8.02
38	II	M	3513	96.93	78.57	50.07	7.69
	I-1	H	1362	95.89	73.94	44.86	7.34
	I-2	H	1502	96.40	74.43	47.07	7.39
39	II	M	3416	95.61	75.26	47.28	7.58
	I-1	H	1599	95.18	75.23	46.90	7.19
	I-2	H	2053	94.93	70.77	44.18	7.11
40	II	M	2827	96.00	73.51	46.06	7.50
	I-1	H	1306	94.56	76.26	44.64	7.96
	I-2	H	1498	93.46	73.70	44.93	7.61
41	II	M	3285	96.99	79.15	49.92	8.55
	I-1	H	1727	93.11	75.16	46.79	8.22
	I-2	H	1406	92.39	71.05	44.52	7.54
42	II	M	2998	96.23	74.45	50.43	11.11
	I-1	H	1208	92.80	72.10	48.76	9.93
	I-2	H	1638	91.58	71.06	47.99	10.74
43	II	M	3262	95.25	71.46	44.67	7.48
	I-1	H	1318	93.17	70.18	42.49	7.28
	I-2	H	1530	92.94	68.82	42.35	7.58
44	II	M	2852	93.41	66.55	40.74	6.59
	I-1	H	1316	93.92	70.82	42.86	6.99
	I-2	H	1438	87.76	63.77	39.15	6.75
45	II	M	3371	95.67	70.54	44.26	7.18
	I-1	H	1532	89.49	65.99	38.77	7.18
	I-2	H	1559	93.91	68.95	42.53	7.44
46	II	M	3184	94.38	73.34	47.80	7.29
	I-1	H	1320	96.67	76.67	47.05	7.80
	I-2	H	1692	98.29	76.42	47.75	8.04
47	II	M	3252	96.92	78.38	48.31	7.35
	I-1	H	1555	94.60	71.83	41.54	6.62
	I-2	H	1536	94.34	73.24	46.16	7.10
48	II	M	3078	96.82	76.67	48.73	7.24
	I-1	H	1249	87.51	69.26	42.75	6.57
	I-2	H	1550	93.81	72.97	45.23	7.10

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
49	II	M	3481	98.65	76.07	48.78	7.38
	I-1	H	1545	97.15	76.25	46.93	7.38
	I-2	H	1439	96.04	75.75	45.45	7.02
50	II	M	3051	97.08	79.78	50.38	7.54
	I-1	H	1375	97.02	75.49	48.87	7.49
	I-2	H	1728	93.69	72.63	44.27	7.29
51	II	M	3343	97.52	79.30	51.99	7.90
	I-1	H	1323	96.37	76.80	48.07	7.18
	I-2	H	1795	95.43	76.10	46.85	7.52
52	II	M	3215	98.16	79.69	51.20	7.74
	I-1	H	1363	92.22	75.79	48.13	7.34
	I-2	H	1499	96.53	77.72	48.37	7.40
53	II	M	3657	96.64	75.42	49.11	7.38
	I-1	H	1528	99.21	76.90	47.38	7.07
	I-2	H	1587	97.35	74.86	47.83	6.81
54	II	M	3554	96.31	77.21	49.24	7.40
	I-1	H	1249	96.08	75.66	47.96	7.53
	I-2	H	1640	93.66	72.56	44.21	7.01
55	II	M	3860	96.48	78.11	49.25	7.10
	I-1	H	1985	93.55	77.18	47.15	6.65
	I-2	H	1927	96.78	77.79	47.90	6.64
56	II	M	3086	94.52	75.79	48.06	7.45
	I-1	H	1359	96.03	73.14	45.11	6.70
	I-2	H	1574	93.58	73.00	43.58	6.80
57	II	M	3260	96.84	78.25	51.26	6.81
	I-1	H	1439	90.55	74.84	47.53	6.25
	I-2	H	2043	95.35	78.32	49.24	6.66
58	II	M	3294	98.82	80.12	51.43	8.65
	I-1	H	1487	96.30	75.99	46.00	7.73
	I-2	H	1793	93.31	75.57	46.07	7.64
59	II	M	3270	96.91	74.19	46.67	8.13
	I-1	H	1454	90.51	68.64	41.27	7.15
	I-2	H	1509	94.23	71.50	44.40	8.15
60	II	M	3634	96.51	73.20	45.40	7.79
	I-1	H	1371	96.64	71.55	41.79	7.08
	I-2	H	1486	93.67	69.99	40.58	7.27

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
61	II	M	3612	94.66	75.08	47.81	8.14
	I-1	H	1487	92.60	73.97	46.40	7.80
	I-2	H	1643	93.67	73.22	44.00	7.55
62	II	M	3648	93.91	72.81	46.66	7.87
	I-1	H	1729	92.02	69.46	42.86	7.63
	I-2	H	1912	91.63	72.38	43.78	7.79
63	II	M	3724	94.33	73.34	46.62	8.11
	I-1	H	1562	95.07	75.35	45.71	8.00
	I-2	H	1536	93.42	70.83	42.19	7.36
64	II	M	3193	95.24	74.41	47.64	8.49
	I-1	H	1378	96.59	78.88	49.49	8.78
	I-2	H	1456	91.62	74.04	42.99	8.04
65	II	M	3371	94.90	74.28	46.75	7.86
	I-1	H	1444	94.11	69.94	42.04	6.79
	I-2	H	1633	93.39	71.34	44.89	7.72
66	II	M	3265	96.39	78.44	49.19	8.67
	I-1	H	1393	95.84	74.30	47.38	8.18
	I-2	H	1713	95.86	73.09	43.32	7.47
67	II	M	3538	95.34	72.81	42.85	6.42
	I-1	H	1424	93.82	67.63	39.61	5.76
	I-2	H	1484	92.05	66.24	38.41	5.86
68	II	M	3781	95.27	73.71	45.68	7.83
	I-1	H	1314	91.02	72.15	43.15	7.23
	I-2	H	1694	96.87	69.24	41.44	7.02
69	II	M	3396	97.35	77.56	49.59	8.60
	I-1	H	1183	89.77	73.37	46.07	8.20
	I-2	H	1489	94.83	69.38	41.64	7.32
70	II	M	3554	94.43	65.81	40.04	6.53
	I-1	H	1570	93.69	64.52	38.09	5.86
	I-2	H	1796	93.10	64.31	36.19	5.62
71	II	M	3688	93.28	72.86	48.07	7.81
	I-1	H	1515	96.44	76.90	48.65	7.92
	I-2	H	1935	94.99	76.12	46.46	7.60
72	II	M	3765	93.55	71.93	45.21	7.22
	I-1	H	1568	95.92	71.94	44.96	7.65
	I-2	H	1693	94.03	72.06	43.24	7.62

TABLE 5 (Cont.)
TEST RESULTS OF SAMPLES PREPARED FOR EXPERIMENT C

Sample No.	Subsample No.	Testing Method	Sample Weight, gm	Percent Passing Sieves			Percent Loss-by-Washing
				3/4-in.	3/8-in.	No. 8	
73	II	M	3429	93.76	73.20	46.84	7.29
	I-1	H	1918	90.41	70.86	43.12	7.19
	I-2	H	1286	87.25	70.14	41.99	6.77
74	II	M	3638	95.38	71.91	44.64	7.26
	I-1	H	1694	89.61	64.64	39.37	6.85
	I-2	H	1336	100.00	71.71	42.66	7.19
75	II	M	3246	95.93	74.12	46.58	7.64
	I-1	H	1321	91.75	67.90	41.48	7.57
	I-2	H	1685	92.64	69.91	43.09	7.66
76	II	M	3326	96.66	75.89	47.17	7.73
	I-1	H	1484	94.41	75.61	43.35	7.08
	I-2	H	1731	93.24	68.46	40.96	6.99
77	II	M	3681	93.64	73.95	46.18	7.61
	I-1	H	1555	92.73	70.87	42.25	7.52
	I-2	H	1298	96.30	75.12	45.45	7.40
78	II	M	3526	96.68	73.34	47.39	7.40
	I-1	H	1403	96.22	72.42	47.04	7.27
	I-2	H	1713	99.01	75.48	47.75	7.36
79	II	M	3265	95.34	71.82	44.26	7.35
	I-1	H	1280	97.19	74.61	44.92	7.34
	I-2	H	1205	95.60	77.18	46.64	7.55
80	II	M	3590	96.24	74.60	48.89	7.74
	I-1	H	1733	99.13	74.50	45.76	8.02
	I-2	H	1694	95.75	76.21	46.58	8.15
81	II	M	3273	95.33	74.18	48.58	8.49
	I-1	H	1589	97.86	73.32	47.26	9.00
	I-2	H	1555	94.47	66.82	43.34	8.04
82	II	M	2905	94.70	72.87	47.85	8.92
	I-1	H	1649	93.33	76.41	50.52	9.95
	I-2	H	1453	93.53	69.30	45.01	8.53
83	II	M	3061	95.98	73.37	46.19	8.07
	I-1	H	1172	94.11	72.27	45.65	8.11
	I-2	H	1637	94.50	73.37	45.51	7.94
84	II	M	3324	95.64	76.59	47.26	8.03
	I-1	H	1320	100.00	74.62	46.21	8.03
	I-2	H	1620	93.09	75.86	45.93	7.72

TABLE 6
IN-PLACE AGGREGATE UNIFORMITIES OF
VARIOUS CONSTRUCTION PROJECTS

Sources	Variances of Percent Passing Sieves			Variance of Percent Loss-by-Washing
	3/4-in.	3/8-in.	No. 8	
Stillman Pit ¹	4.17	9.88	10.76	0.59
Pifke Pit ¹	4.33	4.28	2.36	0.41
Anderson Pit ¹	3.02	4.80	3.49	0.37
M 36021 ²	2.36	9.52	6.66	0.27
FR 23092 ²	1.27	33.95	31.60	0.64
FR 64015 ²	1.77	7.46	5.15	0.36
I 50062 ²	0.18	16.17	26.46	1.35

¹ Information presented in Research Report No. R-1024, MDOT.

² The construction projects are inspected by the newly developed in-place aggregate acceptance sampling plan.

TABLE 7
THE ESTIMATED SPLITTING ERRORS

Sieve Size	Sample Splitter B		Sample Splitter C (Second Stage Splitting)	
	Side 1	Side 2	Side 1	Side 2
3/4-in.	0.1089	-0.0946	-0.0309	0.0291
3/8-in.	0.8866	-0.6279	-0.0044	0.0149
No. 8	1.0033*	-0.7131*	-0.0311	0.0658
Loss-by-Washing	0.0241	-0.0232	0.0186	-0.0231

* The figure is significantly different from zero at the 95 percent confidence level.

TABLE 8
THE ESTIMATED SPLITTING PRECISIONS

Sieve Size	Sample Splitter B		Sample Splitter C			
	Side 1	Side 2	1st Stage Splitting		2nd Stage Splitting	
			Side 1	Side 2	Side 1	Side 2
3/4-in.	1.3289	0.7381	1.1757	0.9744	2.7516	2.4195
3/8-in.	2.8082	1.4046	1.8403	1.5447	2.9029	2.0411
No. 8	2.8002	1.4299	0.7458	0.6014	2.0359	1.8920
Loss-by-Washing	0.0639	0.0340	0.0380	0.0326	0.0592	0.0518

TABLE 9
ESTIMATES OF THE IN-PLACE
AGGREGATE UNIFORMITIES

Sieve Size	Experiment B	Experiment C
3/4-in.	1.1944	1.5422
3/8-in.	8.0235	6.5195
No. 8	8.0649	4.5449
Loss-by-Washing	0.3433	0.3908

ANALYSIS AND CONCLUSIONS

Splitting Ability

The sample splitter is able to split a sample of aggregate into two portions such that their weight ratio is very nearly constant. However, this constant may not be the ratio of the number of openings on the two sides of the splitter.

The estimated splitting errors and precisions in terms of aggregate gradation are presented in Tables 7 and 8 (see Appendices B and C for details). The symbol * in Table 7 means that the figure is significantly different from zero at the 95 percent confidence level (see Appendix B). In this case, we conclude that there is a splitting error relative to the individual splitter precision (Table 8). While we conclude that the aggregate gradations of the two portions obtained by Splitter B are slightly different, the question remains: Are these differences significant relative to the in-place aggregate uniformity? To answer this question, we compare Table 8 with Table 6 and see that the differences could remain significant if in-place aggregate uniformity is high. Since the splitting precision could positively correlate with the in-place aggregate uniformity, we should compare the splitting precision with the in-place aggregate uniformity of the job from which samples were obtained. The estimated in-place aggregate uniformities are presented in Table 9. Now, comparing Table 8 with Table 9, we see that none of the testing errors presented in Table 7 would be significant. Based on the above analysis, we conclude that the aggregate gradations of the two subsamples obtained by a sample splitter are practically the same. Thus, either subsample can be used to represent the original sample for in-place aggregate inspection purposes.

In Appendices B and C, we show that the larger subsample has less splitting error numerically than the smaller subsample. Moreover, the splitting precision on the larger subsample is higher. Thus, the larger subsample is better than the smaller subsample in representing the original sample.

In Experiment C it was shown that the larger the sample, the higher the precision of the sample splitter. The above statement can be verified by comparing the splitting precision of the first stage splitting with that of the second stage splitting (Table 8). We know that, on the average, the sample weight of Experiment B is larger than that of Experiment C. Therefore, the splitting precision of Splitter B should be higher than that of Splitter C; however, this is not the case. This phenomenon can be explained by one or both of the following reasons:

a) the splitting precision is positively correlated with the in-place aggregate uniformity, or

b) the sample splitter with an equal number of openings on two sides has superior splitting ability.

Statement (b) is given numerically in Table 7. The above arguments indicate that the sample splitter with an equal number of openings on two sides is the better tool for reducing sample size.

Testing Ability

Experiment A is a so-called 'Two-Way Completely Randomized Block Design' with vector observations. A multivariate statistical analysis of variance program was used to analyze the data presented in Table 1. The results led us to conclude at the 95 percent confidence level that the sample weight does not affect the testing ability in measuring aggregate gradation. However, the aggregate gradations measured by the hand and mechanical testing methods are significantly different. To expand on this, we present the 95 percent simultaneous confidence intervals for mean differences of gradations measured by the two testing methods (Table 10). We also present the sample variances of the differences of gradations measured by the two testing methods (Table 11). We see from Table 10 that the differences are quite small, but significant with respect to small sample variations presented in Table 11. After comparing Table 11 with Table 6, we conclude that the small differences in Table 10 are not significant with respect to in-place aggregate uniformity. As pointed out in Appendix A, variances presented in Table 11 are the sum of variances of the two testing errors.

TABLE 10
95 PERCENT SIMULTANEOUS CONFIDENCE INTERVALS FOR
MEAN DIFFERENCES OF GRADATIONS MEASURED BY THE HAND
AND MECHANICAL TESTING METHODS (EXPERIMENT A)

Sieve Size	Mean Difference	Confidence Interval		Significance
		Lower Limit	Upper Limit	
3/4-in.	-0.2163	-0.7979	0.3653	No
3/8-in.	0.2549	-0.0149	0.5247	No
No. 4	-0.6178	-0.8086	-0.4269	Yes
No. 8	0.1067	-0.2113	0.4246	No
Loss-by-Washing	-0.4278	-0.7966	-0.0590	Yes

Based on the above analysis, we reach the following conclusion. The testing precisions of the hand and mechanical testing methods are very high. Consequently, the differences presented in Table 10 can be regarded as the true differences between the two testing methods in measuring the aggregate gradation. However, these slight differences are negligible relative to the in-place aggregate uniformity. That is, the hand and mechanical testing methods are practically the same in measuring aggregate gradation.

TABLE 11
 SAMPLE VARIANCES OF THE DIFFERENCES
 OF GRADATIONS MEASURED BY THE HAND
 AND MECHANICAL TESTING METHODS

Sieve Size	Variance
3/4-in.	0.3537
3/8-in.	0.0761
No. 4	0.0381
No. 8	0.1057
Loss-by-Washing	0.1422

The above conclusion was obtained from an experiment using laboratory samples. We would like to know whether the same conclusions can be reached using field samples. Based on the data obtained in Experiment C, we present the 95 percent simultaneous confidence intervals for mean differences of gradations measured by the two testing methods in Table 12. As one can see from Table 12, the two testing methods are again significantly different, relative to sample variations, in measuring aggregate gradation. After comparing the sample variations (Appendix C) with Table 6, we conclude that the two testing methods would still differ significantly in measuring the percent passing the 3/4-in., 3/8-in., and No. 8 sieves for inspection purposes. This conclusion is contrary to the finding of Experiment A. We do not know the reason for this discrepancy; however, we speculate that aggregate might be degraded during the test (shaking process). Table 12 indicates that the percent passing each sieve measured by the mechanical method is larger than that measured by the hand sieve; that is, the mechanical test method degrades aggregate more than the hand test method. This should also have been true in Experiment A. We recall that there are two shaking processes in Experiment A: separating aggregate into

five piles, and testing samples. Degradation should mostly occur on the initial shaking. That is, there is probably no significant degradation during the second shaking. This would explain the discrepancy between Experiment A and Experiment C in measuring percent passing. Since the percent loss-by-washing is measured before shaking, this observation would not apply. This is why Experiments A and C agree on loss-by-washing.

TABLE 12
95 PERCENT SIMULTANEOUS CONFIDENCE INTERVALS FOR
MEAN DIFFERENCES OF GRADATIONS MEASURED BY THE HAND
AND MECHANICAL TESTING METHODS (EXPERIMENT C)

Sieve Size	Mean Difference	Confidence Interval		Significance
		Lower Limit	Upper Limit	
3/4-in.	-1.75	-2.4702	-1.0175	Yes
3/8-in.	-2.51	-3.4181	-1.5953	Yes
No. 8	-3.07	-3.6396	-2.4939	Yes
Loss-by-Washing	-0.34	-0.4721	-0.2088	Yes

If the difference between the two testing methods is aggregate degradation during the test, the Department should use the mechanical test method for the following reasons:

a) samples are obtained from the construction sites prior to compaction; the compaction process degrades aggregate, and

b) the mechanical test method would be subject to less operator variations once this method is standardized.

Remarks

Suppose that each sample is about W grams. The aggregate gradation of this sample can be measured in the following three ways:

1) the whole sample is tested,

2) the sample is split into two subsamples by a sample splitter with equal number of openings on two sides and only one of the two subsamples is tested,

3) as in (2) the sample is split into two subsamples; however, both subsamples are tested and the test results combined as the aggregate gradation of the sample.

The sample variation for each method can be expressed as:

$$\sigma_1^2 \text{ (Method 1)} = \sigma_{P1}^2 + \sigma_t^2 \quad (1)$$

$$\sigma_2^2 \text{ (Method 2)} = \sigma_{P1}^2 + \sigma_e^2 + \sigma_t^2 \quad (2)$$

and

$$\sigma_3^2 \text{ (Method 3)} = \sigma_{P1}^2 + 1/2 \sigma_t^2 \quad (3)$$

where σ_{P1}^2 is the in-place aggregate uniformity measured by samples with weight \bar{W} grams, σ_e^2 is the splitting precision and σ_t^2 is the testing precision.

We concluded in the research study that the testing precision is high and negligible relative to the in-place aggregate uniformity. Thus, method (3) is slightly superior to methods (1) and (2) in terms of sample variation. However, this benefit is achieved by doubling the testing cost. Bearing the cost in mind, method (1) should be used in preference to method (3) unless the sample weight is beyond the capacity of the testing equipment.

When the sample weight is beyond the testing capacity, the sample would have to be split into two subsamples; that is, we must choose method (2) or method (3). If we choose method (3), we will be able to decrease the sample variation by $\sigma_e^2 + 1/2 \sigma_t^2$. Again, this benefit is achieved by doubling the testing cost. The desirability of method (2) depends upon the magnitude of σ_e^2 . In general, σ_e^2 correlates positively with σ_{P1}^2 . Thus, when the in-place aggregate uniformity is high, method (2) would be suitable for measuring aggregate gradation.

Instead of taking oversize samples, we shall take samples of weight, say about $W/2$ grams. Generally speaking, σ_{P2}^2 would be larger than σ_{P1}^2 since the weight of this sample is within the range of the testing capacity and can be tested in entirety. Thus, the sample variation, σ_4^2 , measured by this method is,

$$\sigma_4^2 = \sigma_{P2}^2 + \sigma_t^2 \quad (4)$$

and, consequently,

$$\sigma_2^2 - \sigma_4^2 = \sigma_{P1}^2 - \sigma_{P2}^2 + \sigma_e^2 \quad (5)$$

The relationship among σ_{P1}^2 , σ_{P2}^2 , and σ_e^2 determines the choice between method (2) and this alternative method. This relationship can be established in principle, but is unknown at this time. However, if the in-place aggregate uniformity is high, σ_{P1}^2 and σ_{P2}^2 would be very close and, therefore, this alternative method should be used for inspection purposes.

APPENDIX A

TESTING PRECISIONS OF THE HAND AND
MECHANICAL TESTING METHODS

In this appendix, we present the statistical formulation of Experiment A. Experiment A is a so-called 'Two-Way Completely Randomized Block Design' with vector observations. A multivariate statistical analysis of variance program was then used to analyze the data. The results led us to conclude at the 95 percent confidence level that sample weight does not affect the testing ability in measuring aggregate gradation. However, aggregate gradations measured by the hand and mechanical testing methods are significantly different. To fully explain the above statements, we present the 95 percent simultaneous confidence interval for mean differences of gradations measured by hand and mechanical testing methods in Table 10. We also present the sample variations of the differences of gradations measured by hand and mechanical testing methods in Table 11. We see from Table 10 that the mean gradation differences between the two testing methods are quite small, but statistically significant relative to variations presented in Table 11. The question is, do these differences remain significant relative to the in-place aggregate uniformity? To answer this question, we gather and present the in-place aggregate uniformity of various construction projects in Table 6. We see that sample variations (Table 11) are negligible relative to the in-place aggregate uniformity. Thus, the slight differences between the two testing methods would not be significant relative to the in-place aggregate uniformity. This means that hand and mechanical testing methods are practically the same in measuring aggregate gradation.

Now, we would like to interpret these figures in Table 11. For this purpose, we denote $(X_{11}, X_{12}, X_{13}, X_{14}, X_{15})$ and $(X_{21}, X_{22}, X_{23}, X_{24}, X_{25})$ to be the test results (aggregate gradations) measured by the hand and mechanical testing methods, respectively. Namely, $X_{i1}, X_{i2}, X_{i3},$ and X_{i4} are the percent passing the 3/4-in., 3/8-in., No. 4, and No. 8 sieves, respectively, and X_{i5} is the percent loss-by-washing. We also denote $(T_{11}, T_{12}, T_{13}, T_{14}, T_{15})$ and $(T_{21}, T_{22}, T_{23}, T_{24}, T_{25})$ to be the testing errors due to the hand and mechanical testing methods, respectively. Then, X_{ij} can be expressed by the following equation:

$$X_{ij} = P_j + A_{ij} + T_{ij} \quad (A-1)$$

In Eq. (A-1), P_j is the target value set for making up the sample and A_{ij} is the difference between the target value and the true value of the sample. Now, we define D_j to be the difference between X_{1j} and X_{2j} . Then,

$$D_j = X_{1j} - X_{2j} = (A_{1j} - A_{2j}) + (T_{1j} - T_{2j}) \quad (A-2)$$

If we denote $\mu(X)$ and $\sigma^2(X)$ to be the mean and variance, respectively, of the variable X , then we obtain from Eq. (A-2) the following results,

$$\mu(D_j) = \mu(A_{1j}) - \mu(A_{2j}) + \mu(T_{1j}) - \mu(T_{2j}) \quad (A-3)$$

and

$$\sigma^2(D_j) = \sigma^2(A_{1j} - A_{2j}) + \sigma^2(T_{1j}) + \sigma^2(T_{2j}) \quad (A-4)$$

Viewing our sample preparation procedures, it is reasonable to assume that $\mu(A_{1j}) = \mu(A_{2j})$. Thus, we have

$$\mu(D_j) = \mu(T_{1j}) - \mu(T_{2j}) \quad (A-5)$$

Equation (A-5) indicates that D_j , $j = 1, \dots, 5$, can be used to measure the differences between the two testing methods. It is obvious from Eq. (A-4) that for each i and j ,

$$\sigma^2(T_{ij}) \leq \sigma^2(T_{1j}) + \sigma^2(T_{2j}) \leq \sigma^2(D_j) \quad (A-6)$$

The estimates of $\sigma^2(D_j)$, $j = 1, \dots, 5$, were presented in Table 11. We note that $\sigma^2(T_{1j})$ and $\sigma^2(T_{2j})$, $j = 1, \dots, 5$, are the testing precisions of the hand and mechanical testing methods, respectively. We see from Eq. (A-6) and Tables 6 and 11 that the testing precisions of the hand and mechanical testing methods are very high relative to in-place aggregate uniformity.

APPENDIX B
SPLITTING PRECISION OF THE
GILSON SAMPLE SPLITTER

In this appendix, we present the statistical formulation of Experiment B. For a sample of W grams, we denote W_1 and W_2 to be the weights of Subsamples 1 and 2, respectively. Thus, $W = W_1 + W_2$. Define $R_1 = W_1/W$. If every aggregate grain acts independently, i.e., every aggregate grain has the same freedom to be received by any one of the nine openings, the expected value of R_1 should be $4/9$. Based on the data presented in Table B-1, the sample average \bar{R}_1 and sample $S^2(R_1)$ are 0.423 and 0.00049, respectively. The 95 percent confidence interval for the mean value of R_1 is (0.4158, 0.4303). This interval does not cover the theoretical value $4/9$. This means that this sample splitter does not split a sample into two portions with weight ratio 4:5. Since $S^2(R_1)$ is small, we conclude that Side 1 receives 42.3 percent of the total sample weight.

For each sample, we denote $(W_{i1}, W_{i2}, W_{i3}, W_{i4})$ and $(X_{i1}, X_{i2}, X_{i3}, X_{i4})$ to be the test results of Subsample i , $i = 1, 2$, in terms of weight and percentage, respectively. Namely, $W_{i1}(X_{i1})$, $W_{i2}(X_{i2})$, and $W_{i3}(X_{i3})$ are the total weight (percent) passing the 3/4-in., 3/8-in., and No. 8 sieves, respectively, and $W_{i4}(X_{i4})$ is the total weight (percent) loss-by-washing. The relationship between W_{ij} and X_{ij} is $X_{ij} = 100 W_{ij}/W$. As termed in Appendix A, $(X_{i1}, X_{i2}, X_{i3}, X_{i4})$ is the aggregate gradation of Subsample i . We also denote (P_1, P_2, P_3, P_4) to be the true aggregate gradation of the sample. Then, viewing the sample preparation and testing procedures, W_{ij} can be expressed as the sum of true value, splitting, and testing errors. That is,

$$W_{ij} = \frac{P_j}{100} W_i + E_{ij} + T_{ij} \quad (\text{B-1})$$

In Eq. (B-1), $(E_{i1}, E_{i2}, E_{i3}, E_{i4})$ and $(T_{i1}, T_{i2}, T_{i3}, T_{i4})$ are, respectively, the splitting and testing errors on Subsample i . The units of E_{ij} and T_{ij} are grams. It is clear that $E_{1j} + E_{2j} = 0$ for each j . To facilitate the later analysis, we define the following variables: $X_j = (W_{1j} + W_{2j}) / (W_1 + W_2)$, $e_{ij} = 100 E_{ij}/W_i$ and $t_{ij} = 100 T_{ij}/W_i$. It is clear that (X_1, X_2, X_3, X_4) are the test results (gradation) of the original sample. e_{ij} and t_{ij} are, respectively, the splitting and testing errors on Subsample i in percent. Using the above relations, we obtain from Eq. (B-1) the following basic equation,

$$X_j = P_j + R_1 t_{1j} + (1 - R_1) t_{2j} \quad (\text{B-2})$$

and

$$X_{ij} = P_j + e_{ij} + t_{ij} \quad (\text{B-3})$$

To study the splitting ability of the sample splitter, we define the new variable D_{ij} as the difference of X_{ij} and X_j . That is,

$$D_{1j} = X_{1j} - X_j = e_{1j} + (1 - R_1) (t_{1j} - t_{2j}) \quad (\text{B-4})$$

and

$$D_{2j} = X_{2j} - X_j = e_{2j} - R_1(t_{1j} - t_{2j}) \quad (\text{B-5})$$

Using the fact that $E_{1j} + E_{2j} = 0$ and $W_1 + W_2 = 0$, we obtain from Eqs. (B-4) and (B-5) the relation,

$$D_{1j} = \frac{W_2}{W_1} D_{2j} \quad (\text{B-6})$$

Let us use $\mu(Y)$ and $\sigma^2(Y)$ as the mean and variance of the random variable Y , respectively. Since both subsamples were tested by the same testing method and since (as concluded in Appendix A) sample weight does not affect the testing ability in measuring the gradation, we have $\mu(t_{1j}) = \mu(t_{2j})$. Applying this relation to Eqs. (B-4) and (B-5), we obtain the following equations:

$$\mu(D_{ij}) = \mu(e_{ij}) \quad (\text{B-7})$$

$$\sigma^2(D_{1j}) = \sigma^2(e_{1j}) + (1 - R_1)^2 [\sigma^2(t_{1j}) + \sigma^2(t_{2j})] \quad (\text{B-8})$$

and

$$\sigma^2(D_{2j}) = \sigma^2(e_{2j}) + R_1^2 [\sigma^2(t_{1j}) + \sigma^2(t_{2j})] \quad (\text{B-9})$$

We also obtain from Eq. (B-6) the following two equations:

$$\mu(D_{1j}) = -(W_2/W_1) \mu(D_{2j}) \simeq -1.364\mu(D_{2j}) \quad (\text{B-10})$$

and

$$\sigma^2(D_{1j}) = (W_2/W_1)^2 \sigma^2(D_{2j}) \simeq 1.861 \sigma^2(D_{2j}) \quad (\text{B-11})$$

We see from Eq. (B-7) that $(D_{i1}, D_{i2}, D_{i3}, D_{i4})$ can be used to measure the splitting error. We present the sample mean \bar{D}_{ij} of D_{ij} in Table B-1 and the sample covariance matrix S_i of $(D_{i1}, D_{i2}, D_{i3}, D_{i4})$ in Table B-2. One can see from these two tables that $\bar{D}_{1j} = -1.364 \bar{D}_{2j}$ and $S_1 = 1.861 S_2$, approximately. These results support Eqs. (B-10) and (B-11) and, consequently, imply that the splitting process has less effect on the larger subsample (Subsample 2) than the smaller subsample (Subsample 1). Thus, the larger subsample is a better sample in the sense of representing the original sample.

TABLE B-1
SAMPLE MEAN \bar{D}_{ij} OF D_{ij} , $i = 1, 2$ AND $j = 1, \dots, 4$

j	Sieve Size	\bar{D}_{1j}	\bar{D}_{2j}	$-1.364 \bar{D}_{2j}$
1	3/4-in.	0.1089	-0.0946	0.1290
2	3/8-in.	0.8866	-0.6279	0.8565
3	No. 8	1.0033	-0.7131	0.9727
4	Loss-by-Washing	0.0241	-0.0232	0.0317

We would like to test whether or not the estimated splitting errors presented in Table B-1 are statistically significant (relative to the estimated covariance matrix presented in Table B-2). Again, a multivariate statistical analysis of variance program was used to produce the 95 percent simultaneous confidence intervals for means of D_{ij} , $i = 1, 2$, and $j = 1, \dots, 4$. The results are presented in Table B-3. We note that the only interval in Table B-3 that does not cover the zero is the one that measures the percent passing the No. 8 sieve. We further note that the differences in Table B-1 would not be significant at the 99 percent confidence level. The question now is, do these marginal differences remain significant relative to the in-place aggregate uniformity? To answer this question, we compare the sample variations in Table B-2 with Table 6 and conclude that these differences could remain significant if the in-place aggregate uniformity is very high (i.e., the in-place aggregate gradation does not vary significantly from spot to spot). Since the splitting errors could be correlated with the in-place aggregate uniformity, we shall compare the sample variances presented in Table B-2 with the in-place aggregate uniformity of the job from which samples were obtained. The procedures are presented below.

TABLE B-2
 SAMPLE COVARIANCE MATRIX S_i OF $D_{i1}, D_{i2}, D_{i3}, D_{i4}, i = 1, 2$

Sieve Size Combination	S1		S2		1.861 S2	
	Covariance	Variance	Covariance	Variance	Covariance	Variance
	3/4-in. and 3/4-in.	--	1.3289	--	0.7381	--
3/4-in. and 3/8-in.	0.2546	--	0.1276	--	0.2375	--
3/4-in. and No. 8	0.2136	--	0.1008	--	0.1876	--
3/4-in. and L.B.W.*	0.0381	--	0.0166	--	0.0309	--
3/8-in. and 3/8-in.	--	2.8082	--	1.4046	--	2.6140
3/8-in. and No. 8	1.9568	--	0.9796	--	1.8230	--
3/8-in. and L.B.W.	-0.0113	--	0.0018	--	0.0033	--
No. 8 and No. 8	--	2.8002	--	1.4299	--	2.6610
No. 8 and L.B.W.	-0.0218	--	0.0008	--	0.0015	--
L.B.W. and L.B.W.	--	0.0639	--	0.0340	--	0.0633

* L.B.W. means loss-by-washing.

We obtain from Eq. (B-2) that $\sigma^2(X_j) = \sigma^2(P_j) + R_1^2 \sigma^2(t_{1j}) + (1 - R_1)^2 \sigma^2(t_{2j}) \approx \sigma^2(P_j) + 0.179 \sigma^2(t_{1j}) + 0.333 \sigma^2(t_{2j})$. In the Appendix A, we conclude that the testing precision is very high relative to the in-place aggregate uniformity. That is, the amount, $0.179 \sigma^2(t_{1j}) + 0.333 \sigma^2(t_{2j})$, is negligible relative to $\sigma^2(P_j)$. Thus, we have $\sigma^2(X_j) \approx \sigma^2(P_j)$. The estimates of $\sigma^2(X_j)$, $j = 1, \dots, 4$, are presented in Table B-4. Similarly, we obtain from Eqs. (B-4) and (B-5) that $\sigma^2(D_{ij}) \approx \sigma^2(e_{ij})$. The estimated ratios of $\sigma^2(P_j) + \sigma^2(e_{ij})$ to $\sigma^2(e_{ij})$, $j = 1, \dots, 4$, are also presented in Table B-4. Since the confidence interval for the No. 8 sieve in Table B-3 almost covers zero and the in-place aggregate uniformity is low relative to the precision of the sample splitter, we conclude that the aggregate gradations of two subsamples are identical relative to the in-place aggregate uniformity. That is, either subsample can be used to represent the original sample for inspection purposes.

TABLE B-3
95 PERCENT SIMULTANEOUS CONFIDENCE INTERVALS
FOR MEANS OF D_{ij} , $i = 1, 2$ AND $j = 1, \dots, 4$

Sieve Size	Side 1		Side 2		Significance
	Lower Limit	Upper Limit	Lower Limit	Upper Limit	
3/4-in.	-0.5483	0.7660	-0.5843	0.3952	No
3/8-in.	-0.0687	1.8419	-1.3035	0.0477	No
No. 8	0.0493	1.9572	-1.3947	-0.0314	Yes
Loss-by-Washing	-0.1201	0.1682	-0.1283	0.0819	No

TABLE B-4
ESTIMATES OF THE IN-PLACE AGGREGATE UNIFORMITY

Sieve Size	In-Place Aggregate Uniformity, $\sigma^2(P_j)$	$\frac{\sigma^2(P_j) + \sigma^2(e_{1j})}{\sigma^2(e_{1j})}$
3/4-in.	1.1944	2.11
3/8-in.	8.0235	3.86
No. 8	8.0649	3.88
Loss-by-Washing	0.3433	6.37

Summary

1) The aggregate gradations of the two subsamples are slightly different. The differences are statistically significant relative to the splitting precision; however, these differences are no longer significant relative to the in-place aggregate uniformity. Thus, either subsample can be used to represent the original sample for in-place aggregate inspection purposes.

2) The splitting error of the larger subsample is less than that of the smaller subsample. Therefore, the larger subsample is more representative of the original sample.

APPENDIX C
PRECISIONS OF THE TESTING METHODS
AND THE SAMPLE SPLITTER

In this appendix, we present the statistical formulation of Experiment C. For each example, we denote R_1^* to be the weight ratio of Subsample 1 to the original sample. Similarly, we denote R_1 to be the weight ratio of the Subsample I-1 to Subsample I. Since the two sides of the sample splitter have an equal number of openings, we would like to test the hypothesis that the sample splitter can split a sample into two portions of equal weight. Based on the data presented in Table C-5, the sample means, \bar{R}_1^* and \bar{R}_1 , and sample variances, $S_2^2(R_1^*)$ and $S_2^2(R_1)$, of R_1^* and R_1 are as follows:

$$\bar{R}_1^* = 0.479999, S_2^2(R_1^*) = 0.000616$$

$$\bar{R}_1 = 0.481792, S_2^2(R_1) = 0.001407$$

The above sample information indicates at the 95 percent confidence level that the sample splitter does not split a sample into two portions of equal weight. However, this splitter does split a sample into two portions according to a constant weight proportion, independent of the sample weight.

TABLE C-1
SAMPLE MEAN \bar{D}_{ij} OF D_{ij} , $i = 1, 2$, AND $j = 1, \dots, 4$

(Splitting Subsample I Into Two Subsamples
Tested by the Hand Testing Method)

j	Sieve Size	\bar{D}_{1j}	\bar{D}_{2j}	$-1.076 \bar{D}_{2j}$
1	3/4-in.	-0.03086	0.02904	-0.03125
2	3/8-in.	-0.00430	0.01485	-0.01598
3	No. 8	-0.03110	0.06581	-0.07081
4	Loss-by-Washing	-0.01860	-0.02306	0.02481

Splitting Precision

For each sample, we denote W and (P_1, P_2, P_3, P_4) to be the weight and true gradation of Subsample I, respectively. We further denote W_1 and W_2 to be the respective weights of Subsample I-1 and Subsample I-2. As one can see, this portion of the experiment is the same as the one conducted in Appendix B except that two subsamples of Subsample I were tested using the hand testing method. Thus, Eqs. (B-1) through (B-9), Appendix B, are also the system equations for this portion of the experiment.

TABLE C-2
 SAMPLE COVARIANCE MATRIX S_i OF D_{i1} , D_{i2} , D_{i3} , D_{i4} , $i = 1, 2$
 (Splitting Subsample II into Two Subsamples Tested by the Hand Testing Method)

Sieve Size Combination	S_1		S_2		1.157 S_2	
	Covariance	Variance	Covariance	Variance	Covariance	Variance
	3/4-in. and 3/4-in.	--	2.7516	--	2.4195	--
3/4-in. and 3/8-in.	1.5670	--	1.4619	--	1.6914	--
3/4-in. and No. 8	1.0094	--	0.9527	--	1.1023	--
3/4-in. and L.B.W.*	0.1369	--	0.1269	--	0.1468	--
3/8-in. and 3/8-in.	--	2.9029	--	2.6411	--	3.0558
3/8-in. and No. 8	2.0175	--	1.8705	--	2.1688	--
3/8-in. and L.B.W.	0.2231	--	0.2095	--	0.2424	--
No. 8 and No. 8	--	2.0359	--	1.8920	--	2.1890
No. 8 and L.B.W.	0.1915	--	0.1783	--	0.2063	--
L.B.W. and L.B.W.	--	0.0592	--	0.0518	--	0.0599

* L.B.W. means loss-by-washing.

We note that T_{ij} and t_{ij} in Eqs. (B-1) through (B-9) should be interpreted as the testing errors of the hand testing method. Replacing W_2/W_1 in Eq. (B-6) with the estimated ratio obtained from this experiment, we have

$$\mu(D_{1j}) \simeq -1.076\mu(D_{2j}) \quad (C-1)$$

and

$$\sigma^2(D_{1j}) \simeq 1.157 \sigma^2(D_{2j}) \quad (C-2)$$

The sample information regarding D_{ij} is presented in Tables C-1 through C-3. Again, the results presented in Tables C-1 and C-2 support Eq. (B-6). We see from Table C-3 that every interval contains zero. This leads us to conclude at the 95 percent confidence level that the sample splitter does split a sample into two portions of equal gradation. Comparing Table C-2 with Table 6, we see that the splitting precision is high relative to the in-place aggregate uniformity and, therefore, conclude that either Subsample I-1 or Subsample I-2 can be used to represent Subsample I for in-place aggregate inspection purposes.

TABLE C-3
95 PERCENT SIMULTANEOUS CONFIDENCE INTERVALS
FOR MEANS OF D_{ij} , $i = 1, 2$, AND $j = 1, \dots, 4$
(Splitting Subsample I Into Two Subsamples
Tested by the Hand Testing Method)

Sieve Size	Side 1		Side 2	
	Lower Limit	Upper Limit	Lower Limit	Upper Limit
3/4-in.	-0.6138	0.5521	-0.5176	0.5757
3/8-in.	-0.6031	0.5944	-0.5563	0.5860
No. 8	-0.5326	0.4704	-0.4176	0.5492
Loss-by-Washing	-0.0669	0.1041	-0.1031	0.0570

Testing Precision

Samples used for studying testing precision in Appendix A were made up according to the predetermined target gradations. In this appendix, the actual field samples are used for this purpose.

For each sample, we denote W^* to be the weight of Subsample II. That is, the weight of this sample (the original sample) is $U = W + W^*$. We also denote $(W_1^*, W_2^*, W_3^*, W_4^*)$ and $(X_1^*, X_2^*, X_3^*, X_4^*)$ to be the test results of Subsample II in terms of weight and gradation, respectively, and (Q_1, Q_2, Q_3, Q_4) to be the true gradation of the sample. If we denote F_j^* and T_j^* to be the respective splitting and testing (mechanical testing method) errors on Subsample II and define $f_j^* = 100 F_j/W^*$ and $t_j^* = 100 T_j/W^*$, we have the following familiar equations:

$$W_j^* = \frac{Q_j}{100} W^* + F_j^* + T_j^* \quad (C-3)$$

and

$$X_j^* = Q_j + f_j^* + t_j^* \quad (C-4)$$

Let F_j be the splitting error on Subsample I. Then, $F_j = -F_j^*$. If we define $f_j = 100 F_j/W$, then the relation between P_j and Q_j can be expressed by the following equation.

$$P_j = \frac{Q_j W + 100 F_j}{W} = Q_j + F_j \quad (C-5)$$

Substituting this equation into Eq. (B-2), we have

$$X_j = Q_j + f_j + R_2 t_{1j} + (1 - R_2) t_{2j} \quad (C-6)$$

We note that $R_2 t_{1j} + (1 - R_2) t_{2j}$ is the testing error (hand testing method) on Subsample I and, for simplicity, is denoted as t_j . Thus, Eq. (C-6) can be rewritten as

$$X_j = Q_j + f_j + t_j \quad (C-7)$$

We now form the difference, $H_j = X_j^* - X_j$, to study the testing effects. We obtain from Eqs. (C-4) and (C-7) the following important relation.

$$H_j = f_j^* - f_j + t_j^* - t_j = \left(H \frac{W^*}{W}\right) f_j^* + t_j^* - t_j \quad (C-8)$$

Thus, we obtain from Eq. (C-8) that

$$\sigma^2(H_j) = \left(H \frac{W^*}{W}\right)^2 \sigma^2(f_j^*) + \sigma^2(t_j^*) + \sigma^2(t_j) \quad (C-9)$$

As concluded in Appendix A, $\sigma^2(t_j^*) + \sigma^2(t_j)$ is very small, $\sigma^2(H_j)/(H \frac{W^*}{W})$ can be used to approximate $\sigma^2(f_j)$. The sample information regarding $\sigma^2(H_j)$ and $\sigma^2(f_j^*)$ is presented in Table C-4. Comparing S_f^* in Table C-4 with S_2 in Table C-2 we see that the splitting precision is higher when the original sample is larger. This fact, together with the conclusions presented at the beginning of this section and those of Appendix B allow us to safely assume that the sample splitter is able to divide a sample into two portions of equal gradation, independently of the sample weight. That is, $\mu(f_j^*) = 0$, and, consequently, $\mu(H_j) = \mu(t_j^*) - \mu(t_j)$. This means that (H_1, H_2, H_3, H_4) can be used to measure the difference between the hand and mechanical testing methods in measuring aggregate gradation. The sample information regarding X_j^* , X_j and their difference H_j is presented in Table C-5. The 95 percent simultaneous confidence intervals for $\mu(H_j)$, $j = 1, 2, 3$, and 4, are also presented in Table C-5. As one can see from Table C-5, no interval covers the zero. Thus, we conclude at the 95 percent confidence level that the two testing methods are significantly different on gradation measurements relative to sample variations presented in Table C-4. The question that remains to be answered is whether the differences remain significant relative to the in-place aggregate uniformity. After comparing Table C-6 with Table 10 and Table C-4 with Table 6, we believe that the hand testing method would still not agree with the mechanical testing method in measuring the percent passing the 3/4-in., 3/8-in., and No. 8 sieves relative to the average in-place aggregate uniformity.

TABLE C-4
 SAMPLE COVARIANCE MATRICES S_H AND S_f^* OF
 (H_1, H_2, H_3, H_4) AND $(f_1^*, f_2^*, f_3^*, f_4^*)$

Sieve Size Combination	S_H		S_f^*	
	Covariance	Variance	Covariance	Variance
3/4-in. and 3/4-in.	--	4.2717	--	0.9820
3/4-in. and 3/8-in.	3.2894	--	0.7662	--
3/4-in. and No. 8	1.7578	--	0.4041	--
3/4-in. and L.B.W.*	0.1895	--	0.0436	--
3/8-in. and 3/8-in.	--	6.7255	--	1.5461
3/8-in. and No. 8	3.6017	--	0.8280	--
3/8-in. and L.B.W.	0.4799	--	0.1103	--
No. 8 and No. 8	--	2.6566	--	0.6107
No. 8 and L.B.W.	0.3390	--	0.0779	--
L.B.W. and L.B.W.	--	0.1403	--	0.0323

* L.B.W. means loss-by-washing.

TABLE C-5
 SAMPLE MEANS OF X_j^* , X_j AND H_j AND THE 95 PERCENT
 SIMULTANEOUS CONFIDENCE INTERVALS FOR THE
 MEANS OF H_j , $j = 1, 2, 3, \text{ AND } 4$

Sieve Size	\bar{X}_j^*	\bar{X}_j	\bar{H}_j	Confidence Interval	
				Lower Limit	Upper Limit
3/4-in.	95.81	94.06	1.75	1.0175	2.4702
3/8-in.	74.61	72.10	2.51	1.5953	3.4181
No. 8	47.37	44.30	3.07	2.4939	3.6396
Loss-by-Washing	7.67	7.33	0.34	0.2088	0.4721

For information purposes, we present the procedures for estimating the in-place aggregate uniformity based on the data presented in Table 5. By multiplying $W/100$ to Eq. (C-7), we obtain:

$$W_j = \frac{Q_j}{100} W + F_j + T_j \quad (C-10)$$

The relation between T_j and t_j is that $t_j = 100 T_j/W$. If we define $U_j = 100 (W_j^* + W_j)/(W^* + W)$, then we obtain from Eqs. (C-3) and (C-10)

$$U_j = Q_j + R_1^* t_j^* + (1 - R_1^*) t_j \quad (C-11)$$

Thus,

$$\sigma^2(U_j) = \sigma^2(Q_j) + (R_1^*)^2 \sigma^2(t_j^*) + (1 - R_1^*)^2 \sigma^2(t_j) \quad (C-12)$$

Since the testing precisions are relatively high as concluded in Appendix A, $\sigma^2(U_j)$ can be used to approximate $\sigma^2(Q_j)$ which measures the in-place aggregate uniformity. We present the sample estimates of $\sigma^2(U_j)$, $j = 1, 2, 3, \text{ and } 4$, in Table C-6.

Summary

1) The sample splitter is able to divide a sample into two portions according to a constant weight ratio, independently of the sample weight.

2) The sample splitter is able to divide a sample into two portions of equal gradation.

3) The precision of the sample splitter correlates positively with the sample weight.

4) The two testing methods do not agree with each other in measuring the percent passing the 3/4-in., 3/8-in., and No. 8 sieves. However, these two methods agree in measuring the percent loss-by-washing.

TABLE C-6
ESTIMATES OF IN-PLACE AGGREGATE UNIFORMITY

Sieve Size	Estimate of $\sigma^2(U_j)$
3/4-in.	1.5422
3/8-in.	6.5195
No. 8	4.5449
Loss-by-Washing	0.3908